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It is commonplace of survey research that if anything can possibly go wrong, it will. The collection of data in large scale interview surveys is filtered through a series of intermediaries; the phenomena studied are rarely observed directly. The ultimate source of information, the respondent, is for most purposes only another intermediary, with his own flaws and inadequacies in his ability to transmit accurate reports of the traits, characteristics, events or conditions about which the total stranger we have sent to his door is inquiring, generally for some reason which to the respondent is remote, abstract or irrelevant. He is asked to respond to questions out of context to their real life situation, concerning matters about which good friends, and sometimes even wives, are usually not expected to inquire. After completing the interview, or possibly that night, after putting the children to bed and washing the dishes, the interviewer reviews the questionnaire, to see that each space is appropriately filled in.

At survey headquarters, the one-hundred or so answers on the questionnaire are transferred to a record and manipulated to reduce them to the five, ten, twenty, or fifty dimensions which constitute the data of the survey. The data are then further manipulated to produce a set of tables and numbers which presumably have some bearing on the problem which prompted the study. Considering the multiple possibilities of error, it is reasonable and necessary to ask how good are the data we have collected.

Up to the point at which the interviewer knocked at the door, we had a precise estimate of the range of error expected from sampling, assuming an acceptable probability sampling design consistent with the objectives of the survey. At the respondent's door, however, errors attributable to the interviewer, the respondent or the interviewer-respondent interaction enter into the measurement. Regardless of the precision expected from the sample design, we have not accounted for the accuracy of the measurement until we have some notion of the kind and extent of the response errors entering into the estimates.

In this paper we will present data on the magnitude and type of response error and response bias in the reporting of social security benefit income and we will examine some of the characteristics associated with response bias. The data are based on a national area probability sample of all persons aged 62 and over, the 1963 Survey of the Aged, conducted by the Social Security Administration, in cooperation with the Bureau of the Census. Validation of the benefit income was obtained from a complete matching of the survey interviews with the benefit income data from Social Security record files. The concern with response validity predates the sample survey by many years, as illustrated by the laws of perjury, rules of evidence and trial cross-examination procedures. Only in recent years has a systematic body of data and of theory on response error begun to develop.1/ Unlike sampling error, however, there is little existing theory as yet on which a priori judgments or estimates of response error may be based.

One of the difficulties in the study of response error is the limited availability of criterion data for assessment of the survey data. In many cases, the available criteria contain similar response errors, or a different but equally unknown set of estimation errors. In some areas of measurement, such as personality and attitudes, the existence of an external reality against which the data can be measured may itself be in question.

Most of the things with which we are concerned in survey research, however, do have an external reality or may be represented by events, acts, or conditions external to the reporting unit.

The 1963 Survey of the Aged provided an unusual opportunity to study response error in income reporting through the virtually complete matching of survey data from an area probability sample with a large record system, the Social Security benefit and earnings records. The survey data were collected in two stages and included information to identify respondents with the records and information on specific income sources, including benefit income and earnings. Interviews were completed with 7,500 survey units, consisting of a person aged 62 and over and his spouse, if any. Matching record information was located for all but 2 cases per 1,000 interviewed.2/ The benefit income reporting comparison is based on the total of 4,727 beneficiary units, representing ll million couples and aged individuals. The sample is selfweighting, but has not been adjusted for population and non-interview factors in this analysis.

The matching operation was primarily undertaken to insure accurate reporting of beneficiary status and benefit income. The methodological analysis of benefit income reporting was a byproduct of the study rather than a major objective. Both benefit income amounts were kept in the data tape, but only the "true" or benefit record value was reported in the study, with the exception of a \$60 range of error allowed to reduce the review and editing work load.3/

The accuracy of the respondent's income estimate was measured from the discrepancies between the amount recorded in the benefit record, the "true" value for the respondent, and the amount obtained in the interview. The mean benefit income reported is the survey estimate of the "true population value," without response error. The differences between the Social Security benefit record and the interview data represent response error, including both variable error and bias.

In order to examine some of the possible sources of response error, the data were analyzed by age, type of unit, education, income thirds, living arrangements, and amount of OASDI benefit income. Beneficiaries who did not correctly report their beneficiary status in the interview are included in the analysis.

Data from the record match

1. Response error and bias

Table 1 shows the distribution of reporting differences between the SSA benefit record and the interview for all beneficiary units, couples and the nonmarried. Almost half of the beneficiary income reports matched exactly with the benefit record and more than threefourths of the beneficiaries were within a range of plus or minus \$60. About one-fourth of the responses were distributed among over and underestimates of more than \$60, but the proportion with underestimates was \$ percent higher. Most of the net differences in under and over reporting were large amounts in excess of \$300.

The net income understatement or bias was \$61, about 6 percent of the Mean Total Benefit Income (TBI) of \$1,052, with a standard error of approximately \$5. As table 2 indicates, almost all of the bias was attributable to the small proportion of units with large under reports: four-fifths of the bias was from errors in excess of \$500. Response errors of less than \$301 were almost entirely random variation and made little or no contribution to the bias.

Considering that more than half of the aged received less than \$1,000 in benefits and only one in 20 received more than \$2,000, errors in excess of \$300 are quite sizable for this group and would seem to be more than simple inability to recall or work out an approximate total. Faulty reporting of beneficiary status may account for as much as a third of this; about two percent of the beneficiaries did not report their beneficiary status, but the variation in the proportion not reporting did not show any direct relationship to the proportion with large underreporting errors.

Other factors related to the particular conditions of benefit payments and to the questionnaire method may also have resulted in the omission of income amounts. Large back payments for delayed benefit awards, benefit terminations and suspensions may result in unusual or irregular payments. The questionnaire was set up primarily to obtain accurate totals on regular income and may thus have led to some understatement of irregular payments. Our data do not examine this possibility directly, but there is some support for this in the differences in the size and error rate of the bias by age groups.

2. Matching reports under \$61

We were also interested in the effect on the income distribution and mean of the editing procedure allowing differences of up to \$60 to be accepted as matching amounts. This was done in order to reduce the edit rejection rate and to eliminate extensive clerical checking for small differences.

As shown in tables 1 and 2, the \$60 allowance was an effective procedure for reducing editing at little or no cost in bias. Reporting differences under \$61 were randomly distributed between under and over reporting; 31 percent of the respondents had reporting errors under \$61, but only 2 percent more under reported than over reported income. The net effect on the total benefit income was negligible, an average understatement of \$1.

From an examination of the distribution, it would seem that our allowance was overly conservative and that we could have accepted all matching differences up to three hundred without affecting the mean Total Benefit Income (TBI) or introducing a serious distortion of the income distribution.

3. Sources of selective bias

The several variables selected for analysis were chosen on the basis of experience with or assumptions about their relationship to response validity. Differences between the Current Population Survey and Office of Business Economics estimates of the number of low-income families have been attributed, in part, to transfer income, including social security benefits.4/ Socio-economic factors such as education and income have frequently been found to be positively associated with accuracy of reporting.5/ Age, as an indicator of mental and physiological responsiveness should be negatively associated with validity, while living arrangements, as a measure of independent functioning, should have a positive relationship.

The results of this examination are summarized in table 3, for total beneficiary units. The relationship of these variables to response validity is examined further in tables 4 and 5, with marital status, sex and age held constant.

With the exception of living arrangements, none of the assumptions on the relationship between income reporting and characteristics were supported by the data on total beneficiaries. The younger beneficiaries have a greater net error than the beneficiaries aged 73 and over; the upper-income, larger benefit and higher education groups have greater net error rates than the lower income and education groups. The absolute differences, however, were relatively small and are not statistically significant for the Education and Income characteristics. Respondents with "no answer" on income, however, do have a significantly larger response bias.

The response error rates for beneficiary unit types, in table 4, are generally consistent with those for all beneficiary units, and provide little evidence of selective bias in the reporting of benefit income. Only in living arrangements and age are the percent bias relationships consistent and substantial.

In table 5, age appears to account for most of the bias variation in the benefit income groups, with the 62-64 year age group having biases two to three times that of the older age groups. This is also the age group most likely to have irregular payments because of recent benefit awards and interruptions in payments.

Respondents whose total income was unreported also tended to have among the highest rates of bias, even when age and marital status were controlled.

4. Effect of response error on the standard error

The survey estimate of the standard error of the mean, 341, includes sampling variability, but no response variability. Without the adjustment for bias, from the benefit record, the apparent standard error would have been approximately the same or slightly smaller than the standard error of the record data mean. The mean square error, or "true standard error," however, would have been almost twice that of the apparent standard error if there had been no correction for bias:

Survey estimate of standard error

of mean corrected for bias	<u>×</u> 841
Standard error of bias	ē 5
True standard error of reported	
mean (mean square error) \sqrt{M}	ISE 74*
(m F (F ()) (F F ()) 2)	

 $(MSE = \overline{x} + \overline{e} + 2r \overline{x} e + (Bias)^2)$

* Assuming the coefficient of correlation, $r \leq 1.5$

DISCUSSION

On the basis of these data, we can say with reasonable confidence that the reporting of Social Security benefit income is relatively accurate in surveys of the aged. With few exceptions, respondents reported benefit income with great accuracy. The bias, or net error was small and was confined to understatements by 5 to 7 percent of the reporting units, rather than to constant error or consistent under-reporting. The bias, moreover, was relatively consistent among education and income groups and was selectively distributed to a marked degree only by age.

The extent of understatement of benefit income was also consistent with the findings of a variety of other income studies. The 1950 Census underestimated the National Income accounts by 9 percent.6/ As Guthrie reported last year, the 1960 Consumer Expenditures study underestimated income by 6 percent and the Michigan Consumer Finance Surveys ranged from 3 percent to 13 percent in income underestimates during 1947-1955.7/

Aside from the immediate purposes of the Aged Survey, the data also have some general implications for methods of controlling response bias. The data strongly suggest that income reporting is not a problem of the sensitivity of the instrument. There was little tendency towards "yard-stick" error or consistent individual under-reporting of benefit income. The bias was due almost entirely to a small proportion of cases with gross under-statements. It was also associated with incomplete reporting in other financial areas. A good part of this would be accounted for by people who did not report beneficiary status, probably because of suspended payments and other changes in their beneficiary status. Others may have omitted lump-sum and retroactive payments.

As Ferber concluded from his study of time deposits, it is doubtful that methods designed to increase the accuracy of response overall would be effective for those income reports. Most respondents are already providing adequate answers. The deviant case is the one which presents the problem.

One possibility would be the use of questions directed specifically towards recall of unusual or one-time payments. A good part of the net error came from those respondents who did not answer other income questions, however, and it is doubtful that much more could be obtained from them through the refinement of questioning techniques.

Adjustment procedures which take into account not only the demographic characteristics of non-respondents, but also non-response as an attribute may provide more effective means of accounting for response bias. This will, of course, require more knowledge of the special attributes of non-responders and more validation research. Non-responders on total income, for example, had a mean benefit income equal to the high and middle income thirds. Adjustments for non-response which take this into account as an attribute should, therefore, produce better estimates.

The extent to which these findings may be generalized to other questions and other population can, of course, only be determined by further study. Our research plans include a variety of validation studies related to Social Security record data. We are, for example, preparing a similar analysis for earnings from the 1963 Aged Survey. Data on earnings and benefit income reporting will be available for the under age 65 disabled population, from the 1966 Survey of Disabled Adults. Comparison data on the reliability of diagnostic information will also be developed. The proposed Longitudinal Survey Retirement should provide data on the temporal nature of response error.

The developing computer technology and access to large-scale data systems are extending the possibilities for validation studies and much more research may now be done on response error. It is to be hoped that the growing body of research will lead to a more standard treatment of response estimation procedures and that in the future these procedures will be made available in at least as much detail as is now published for sampling procedures.

References

1/ See, for example, William G. Madow, "On Some Aspects of Response Error Measurement," 1965 Proceedings of the Social Statistics Section, American Statistical Association, Washington, D.C. pp. 182-192, 1965; John B. Lansing, Gerald P. Ginsburg and Kaisa Braaten, An Investigation of Response Error, Studies in Consumer Savings No. 2, 1961, pp. 188-204; and William G. Cochran, Sampling Techniques, 2nd ed., John Wiley and Sons, New York, 1963, pp. 374-389.

- 2/ The procedures for record matching are described by Lawrence D. Haber, "Methodological Analyses in the 1963 Survey of the Aged," paper presented at the Gerontological Society meeting, October 1964, Minneopolis, Minnesota; the sample design is summarized in "Technical Note on Source and Reliability of the Estimates for the 1963 Survey of the Aged," Social Security Bulletin, July 1964, pp. 26-28.
- 3/ The benefit income reported in the study also included minor editing adjustments to resolve differences arising from exclusion of children's benefits and changes in marital status.
- 4/ Selma F. Goldsmith, "Low-Income Families and Measures of Income Inequality," Review of Social Economy, 20:1 (March 1962), pp. 1-19; and "The Relation of Census Income Distribution Statistics to Other Income Data," An Appraisal of the 1950 Census Income Data, vol. 13, Studies in Income and Wealth, National Bureau of Economic Research, Princeton, 1958, pp. 70-83.
- 5/ See, for example, Robert Ferber, "The Reliability of Consumer Surveys of Financial Holdings: Time Deposits," J. American Statistical Association, 60:309 (March 1965), pp. 148-163; and Lansing, op.cit., pp.180-181.
- 6/ Herman P. Miller, <u>Income of the American</u> <u>People</u>, John Wiley & Sons, New York, 1955, <u>p. 15</u>.
- <u>7</u>/ Harold W. Guthrie, "Some Methodological Issues in Validation Studies," <u>1965 Proceeding</u>, op. <u>cit.</u>, pp. 193-196.

Persysting differences	Bereficiary units									
(SSA - Interviev)		Married	Honmarried							
(00,1 11001120,1)	10021	couples	Men	Women						
Sample N Percent	4,727 100	2,137 100	713 100	1,877 100						
Under-reporting \$500 or more 301-500 101-300 61-100	6 3 4 2	8 3 5 2	4 2 3 2	6 3 4 2						
Matching report (+560) Exact 11-60	77 46 31	73 39 3 ¹ +	82 7 ¹ + 8	79 44 3 5						
Over-reporting 61-100 101-300 301-500 501 or more	2 4 1 1	2 6 1 1	2 4 1 1	2 2 1 1						
Percent of units Net response error (NRE) (\$61 or more)	8	8	3	9						
Mean (dollars) Total benefit income (TBI) Net under-reported Percent of TBI	\$1,052 \$61 5.8	\$1,351 \$82 6.1	\$911 \$29 3. 2	\$764 \$50 6.6						

Table 1.-- Distribution of response errors for OASDI benefit income by type of beneficiary unit: Percent of beneficiary units aged 62 and over

Table 2.--Percent distribution of response errors Units and aggregate dollars

	Total beneficiary units 52 and over										
Reporting	Percent c	Net under-reporte dollars									
(SSA- Interview)	Fotal	Net Percent Total under-reporting of mean									
	over-reporting)	over-reporting)	TBI								
Base Percent	4,727 100	4,727 100	\$1,052 100	\$61 100							
Size of difference											
\$501+ 301-500 101-300 61-100 1-60 None	7 4 8 4 31 46	5 2 1 * 2 	4.8 0.8 0.2 0.0 0.1	82 14 2 * 2							

* Less than one-half of one percent.

	Number	Mean Totel	Net Under-report				
Selected Characteristics	of Units	Benefit Income	Mean Bies	Percent of TBI			
Total (62 & over)	4,727	\$1,052	\$61	5.8			
Age							
62-64. 65-72. 73 & over.	449 2,309 1,969	834 1,095 1,050	99 71 41	11.9 6.5 3.9			
Living Arrangements							
In household-no relatives In household-with relatives Institutionalized	2,969 1,702 56	1,099 1,980 717	47 82 194	4.3 8.4 27.0			
Education							
Less than 9 years. 9-11 years. 12 years or more. NA.	2,884 556 979 3 08	1,049 1,060 1,106 892	58 47 73 79	5.6 4.5 6.6 8.8			
OASDI Benefit Income							
\$0-499. 500-9999. 1,000-1,499. 1,500-1,999. 2,000-2,499. 2,500 & over.	755 1,752 1,364 568 244 44	416 765 1,220 1,769 2,133 2,885	2 35 68 96 163 916	0.5 4.6 5.5 5.4 7.7 31.7			
<u>Income Terciles a</u> /(65 & over)	4,278	\$1,075	\$57	5.3			
Lowest Middle. Highest. NA.	1,072 1,401 1,323 480	898 1,135 1,132 1,136	29 46 54 165	3.2 4.0 4.7 14.5			

Table 3.--Mean Total Benefit Income and Response Bias: Total Beneficiary Units

a/ Terciles were defined for each array of unit types separately and represent the position of the unit among couples, non-married men and non-married women.

	พบท	ber of ur	nits	Mean total	benefit :	Income(TBI)	Mean ne	t under-re	eported	Percent of TBI under-reported			
Specified characteristics	Married	Nonmarried		Married	Non	married	Married	Nonne	urried	Married	Nonwirried		
••••••••••••••••••••••••••••••••••••••	couples	Men	Women	couples	Men	Women	couples	Men	Women	couples	Ken	Women	
Total (62 and over)	2,137	713	1,877	\$1,351	\$911	\$764	\$82	\$29	\$50	6.1	3.2	6.6	
<u>Δεε</u> 62-64. 65-72 73 and over	204 1,114 819	42 296 375	203 899 775	991 1,357 1,433	** 978 878	696 809 729	166 95 43	** 44 15	42 51 51	16.8 7.0 3.0	** 4.5 1.7	6.1 6.4 7.0	
Living errangements In household-no relatives. In household-with relatives In Institution	1,559 575 3	401 292 20	1,009 835 33	1,355 1,346 **	921 909 **	775 753 **	61 137 **	17 39 **	37 60 **	4.5 10.2 **	1.8 4.3 **	4.7 7.9 **	
Education Less than 9th grade 9-11th grade 12th grade or more NA	1,383 248 417 89	465 71 104 73	1,036 237 458 146	1,333 1,343 1,453 1,193	894 980 991 842	739 788 816 734	73 70 126 54	21 36 20 85	56 27 37 91	5.5 5.2 8.6 4.5	2.3 3.7 2.0 10.1	7.5 3.4 4.5 12.4	
OASDI Benefit income \$00-499. \$00-999. 1,000-1,499. 1,500-1,999. 2,000-2,499. 2,500 or more.	154 467 676 557 240 43	129 290 286 6 1 1	472 995 402 5 3	351 764 1,264 1,771 2,133 **	431 784 1,230 ** **	433 760 1,139 • ** **	31 38 56 86 157 **	17 17 34 ** **	8 39 112 ** **	-9.0 5.0 4.4 4.9 7.3 **	3.9 2.2 2.7 ** **	2.0 5.1 9.8 ** **	
Income terciles (Total <u>65 and over</u>). Lowest. Middle. Highest. NA.	<u>1,933</u> 568 615 509 241	671 168 241 213 49	1,674 336 545 603 190	1,390 1,165 1,528 1,465 1,406	922 678 938 1,088 **	772 555 780 866 837	73 28 68 75 186	28 22 20 24 **	51 34 32 46 153	5.3 2.4 4.5 5.2 13.2	3.0 3.2 2.2 2.2 **	6.6 6.1 4.1 5.3 18.3	

Table 4Mcan total	benefit income	and response bias	s by type of	beneficiary unit
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** Mean and percents not computed for bases of less than 50 cases.

	Nuz	ber of unit	8	Mean total	benefit in	ncome(TBI)	Mean	net under-r	eported	Percent of TBI under-reported Age of head			
Specified characteristics		Age of head			Age of hea	đ		Age of hea	d				
	62-64	65-72	73 and over	62-64	65-72	73 and over	62-64	65-72	73 and over	62-64	65-72	73 and over	
Total (62 and over)	449	2,306	1,969	\$834	\$1,095	\$1,050	\$99	\$71	\$41	11.9	6.5	3.9	
Living arrangements In household-no relatives. In household-with relatives In institution	274 173 2	1,460 834 15	1,235 695 39	800 887 **	1,133 1,035 **	1,125 937 **	65 154 **	56 94 **	31 50 **	8.1 17.4 **	5.0 9.1 **	2.8 5.3 **	
Education Loss than 9th grade 9-11th grade 12th grade or more NA	268 68 90 23	1,415 302 474 118	1,201 186 415 167	836 877 835 **	1,103 1,096 1,100 983	1,032 1,069 1,171 858	90 109 157 **	71 51 69 143	36 19 59 51	10.8 12.4 18.7 **	6.4 4.6 6.2 14.5	3.5 1.8 5.1 6.0	
OA:DI Benefit income \$C-499. 500-999. 1,000-1,499. 2,000-2,499. 2,000-2,499. 2,500 or more.	137 166 110 25 6 5	314 801 748 276 137 33	304 785 506 267 101 6	324 749 1,180 ** **	402 767 1,223 1,771 2,132 **	472 767 1,225 1,766 2,130 **	-23 70 113 ** **	7 29 79 118 189 **	8 34 41 46 125 **	-7.1 9.3 9.6 ** **	1.8 3.8 6.4 6.7 8.9 **	1.6 4.4 3.4 2.6 5.9 **	
Income terciles Lowest Middle Highest NA		449 756 840 264	623 645 485 216		911 1,181 1,106 1,130	888 1,082 1,176 1,142	 	49 51 75 158	14 40 16 173		5.3 4.3 6.8 14.0	1.6 3.7 1.4 15.1	

Table 5.	Mean	total	benefit	income	and	response	bias	ЪУ	selected	characteristics	and	age	of	head	of	beneficiary u	nit
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** Mcan and percent not computed for bases of less than 50 cases.